

LP coefficients for each segment of the data. The output is reconstructed by driving these coefficients with initial conditions consisting of all zeros except for one nonzero spike. A speech synthesizer chip may have of order 10 LP coefficients, which change perhaps 20 to 50 times per second.

- Some people believe that it is interesting to analyze a signal by LPC, even when the residuals x_i are *not* small. The x_i 's are then interpreted as the underlying "input signal" which, when filtered through the all-poles filter defined by the LP coefficients (see §13.7), produces the observed "output signal." LPC reveals simultaneously, it is said, the nature of the filter *and* the particular input that is driving it. We are skeptical of these applications; the literature, however, is full of extravagant claims.

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- Cressie, N. 1991, in *Spatial Statistics and Digital Image Analysis* (Washington: National Academy Press). [3]
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13.7 Power Spectrum Estimation by the Maximum Entropy (All Poles) Method

The FFT is not the only way to estimate the power spectrum of a process, nor is it necessarily the best way for all purposes. To see how one might devise another method, let us enlarge our view for a moment, so that it includes not only real frequencies in the Nyquist interval $-f_c < f < f_c$, but also the entire complex frequency plane. From that vantage point, let us transform the complex f -plane to a new plane, called the z -transform plane or z -plane, by the relation

$$z \equiv e^{2\pi i f \Delta} \quad (13.7.1)$$

where Δ is, as usual, the sampling interval in the time domain. Notice that the Nyquist interval on the real axis of the f -plane maps one-to-one onto the unit circle in the complex z -plane.

If we now compare (13.7.1) to equations (13.4.4) and (13.4.6), we see that the FFT power spectrum estimate (13.4.5) for any real sampled function $c_k \equiv c(t_k)$ can be written, except for normalization convention, as

$$P(f) = \left| \sum_{k=-N/2}^{N/2-1} c_k z^k \right|^2 \quad (13.7.2)$$

Of course, (13.7.2) is not the *true* power spectrum of the underlying function $c(t)$, but only an estimate. We can see in two related ways why the estimate is not likely to be exact. First, in the time domain, the estimate is based on only a finite range of the function $c(t)$ which may, for all we know, have continued from $t = -\infty$ to ∞ . Second, in the z -plane of equation (13.7.2), the finite Laurent series offers, in general, only an approximation to a general analytic function of z . In fact, a formal expression for representing "true" power spectra (up to normalization) is

$$P(f) = \left| \sum_{k=-\infty}^{\infty} c_k z^k \right|^2 \quad (13.7.3)$$

This is an infinite Laurent series which depends on an infinite number of values c_k . Equation (13.7.2) is just one kind of analytic approximation to the analytic function of z represented by (13.7.3); the kind, in fact, that is implicit in the use of FFTs to estimate power spectra by periodogram methods. It goes under several names, including *direct method*, *all-zero model*, and *moving average (MA) model*. The term “all-zero” in particular refers to the fact that the model spectrum can have zeros in the z -plane, but not poles.

If we look at the problem of approximating (13.7.3) more generally it seems clear that we could do a better job with a rational function, one with a series of type (13.7.2) in both the numerator and the denominator. Less obviously, it turns out that there are some advantages in an approximation whose free parameters all lie in the *denominator*, namely,

$$P(f) \approx \frac{1}{\left| \sum_{k=-M/2}^{M/2} b_k z^k \right|^2} = \frac{a_0}{\left| 1 + \sum_{k=1}^M a_k z^k \right|^2} \quad (13.7.4)$$

Here the second equality brings in a new set of coefficients a_k 's, which can be determined from the b_k 's using the fact that z lies on the unit circle. The b_k 's can be thought of as being determined by the condition that power series expansion of (13.7.4) agree with the first $M + 1$ terms of (13.7.3). In practice, as we shall see, one determines the b_k 's or a_k 's by another method.

The differences between the approximations (13.7.2) and (13.7.4) are not just cosmetic. They are approximations with very different character. Most notable is the fact that (13.7.4) can have *poles*, corresponding to infinite power spectral density, on the unit z -circle, i.e., at real frequencies in the Nyquist interval. Such poles can provide an accurate representation for underlying power spectra that have sharp, discrete “lines” or delta-functions. By contrast, (13.7.2) can have only zeros, not poles, at real frequencies in the Nyquist interval, and must thus attempt to fit sharp spectral features with, essentially, a polynomial. The approximation (13.7.4) goes under several names: *all-poles model*, *maximum entropy method (MEM)*, *autoregressive model (AR)*. We need only find out how to compute the coefficients a_0 and the a_k 's from a data set, so that we can actually use (13.7.4) to obtain spectral estimates.

A pleasant surprise is that we already know how! Look at equation (13.6.11) for linear prediction. Compare it with linear filter equations (13.5.1) and (13.5.2), and you will see that, viewed as a filter that takes input x 's into output y 's, linear prediction has a filter function

$$\mathcal{H}(f) = \frac{1}{1 - \sum_{j=1}^N d_j z^j} \quad (13.7.5)$$

Thus, the power spectrum of the y 's should be equal to the power spectrum of the x 's multiplied by $|\mathcal{H}(f)|^2$. Now let us think about what the spectrum of the input x 's is, when they are residual discrepancies from linear prediction. Although we will not prove it formally, it is intuitively believable that the x 's are independently random and therefore have a flat (white noise) spectrum. (Roughly speaking, any residual correlations left in the x 's would have allowed a more accurate linear prediction, and would have been removed.) The overall normalization of this flat spectrum is just the mean square amplitude of the x 's. But this is exactly the quantity computed in equation (13.6.13) and returned by the routine `memcof` as `xms`. Thus, the coefficients a_0 and a_k in equation (13.7.4) are related to the LP coefficients returned by `memcof` simply by

$$a_0 = \text{xms} \quad a_k = -d(k), \quad k = 1, \dots, M \quad (13.7.6)$$

There is also another way to describe the relation between the a_k 's and the autocorrelation components ϕ_k . The Wiener-Khinchin theorem (12.0.12) says that the Fourier transform of the autocorrelation is equal to the power spectrum. In z -transform language, this Fourier transform is just a Laurent series in z . The equation that is to be satisfied by the coefficients in equation (13.7.4) is thus

$$\frac{a_0}{\left| 1 + \sum_{k=1}^M a_k z^k \right|^2} \approx \sum_{j=-M}^M \phi_j z^j \quad (13.7.7)$$

The approximately equal sign in (13.7.7) has a somewhat special interpretation. It means that the series expansion of the left-hand side is supposed to agree with the right-hand side term by term from z^{-M} to z^M . Outside this range of terms, the right-hand side is obviously zero, while the left-hand side will still have nonzero terms. Notice that M , the number of coefficients in the approximation on the left-hand side, can be any integer up to N , the total number of autocorrelations available. (In practice, one often chooses M much smaller than N .) M is called the *order* or *number of poles* of the approximation.

Whatever the chosen value of M , the series expansion of the left-hand side of (13.7.7) defines a certain sort of *extrapolation* of the autocorrelation function to lags larger than M , in fact even to lags larger than N , i.e., *larger than the run of data can actually measure*. It turns out that this particular extrapolation can be shown to have, among all possible extrapolations, the maximum *entropy* in a definable information-theoretic sense. Hence the name *maximum entropy method*, or MEM. The maximum entropy property has caused MEM to acquire a certain “cult” popularity; one sometimes hears that it gives an intrinsically “better” estimate than is given by other methods. Don’t believe it. MEM has the very cute property of being able to fit sharp spectral features, but there is nothing else magical about its power spectrum estimates.

The operations count in `memcof` scales as the product of N (the number of data points) and M (the desired order of the MEM approximation). If M were chosen to be as large as N , then the method would be much slower than the $N \log N$ FFT methods of the previous section. In practice, however, one usually wants to limit the order (or number of poles) of the MEM approximation to a few times the number of sharp spectral features that one desires it to fit. With this restricted number of poles, the method will smooth the spectrum somewhat, but this is often a desirable property. While exact values depend on the application, one might take $M = 10$ or 20 or 50 for $N = 1000$ or 10000 . In that case MEM estimation is not much slower than FFT estimation.

We feel obliged to warn you that `memcof` can be a bit quirky at times. If the number of poles or number of data points is too large, roundoff error can be a problem, even in double precision. With “peaky” data (i.e., data with extremely sharp spectral features), the algorithm may suggest split peaks even at modest orders, and the peaks may shift with the phase of the sine wave. Also, with noisy input functions, if you choose too high an order, you will find spurious peaks galore! Some experts recommend the use of this algorithm in conjunction with more conservative methods, like periodograms, to help choose the correct model order, and to avoid getting too fooled by spurious spectral features. MEM can be finicky, but it can also do remarkable things. We recommend that you try it out, cautiously, on your own problems. We now turn to the evaluation of the MEM spectral estimate from its coefficients.

The MEM estimation (13.7.4) is a function of continuously varying frequency f . There is no special significance to specific equally spaced frequencies as there was in the FFT case. In fact, since the MEM estimate may have very sharp spectral features, one wants to be able to evaluate it on a very fine mesh near to those features, but perhaps only more coarsely farther away from them. Here is a subroutine which, given the coefficients already computed, evaluates (13.7.4) and returns the estimated power spectrum as a function of $f\Delta$ (the frequency times the sampling interval). Of course, $f\Delta$ should lie in the Nyquist range between $-1/2$ and $1/2$.

```
FUNCTION evlmem(fdt,d,m,xms)
INTEGER m
REAL evlmem,fdt,xms,d(m)
    Given d, m, xms as returned by memcof, this function returns the power spectrum estimate
    P(f) as a function of fdt = fΔ.
INTEGER i
REAL sumi,sumr
DOUBLE PRECISION theta,wi,wpi,wpr,wr,wtemp      Trigonometric recurrences in double
theta=6.28318530717959d0*fdt                    precision.
wpr=cos(theta)                                  Set up for recurrence relations.
wpi=sin(theta)
wr=1.d0
wi=0.d0
sumr=1.                                          These will accumulate the denominator of (13.7.4).
```

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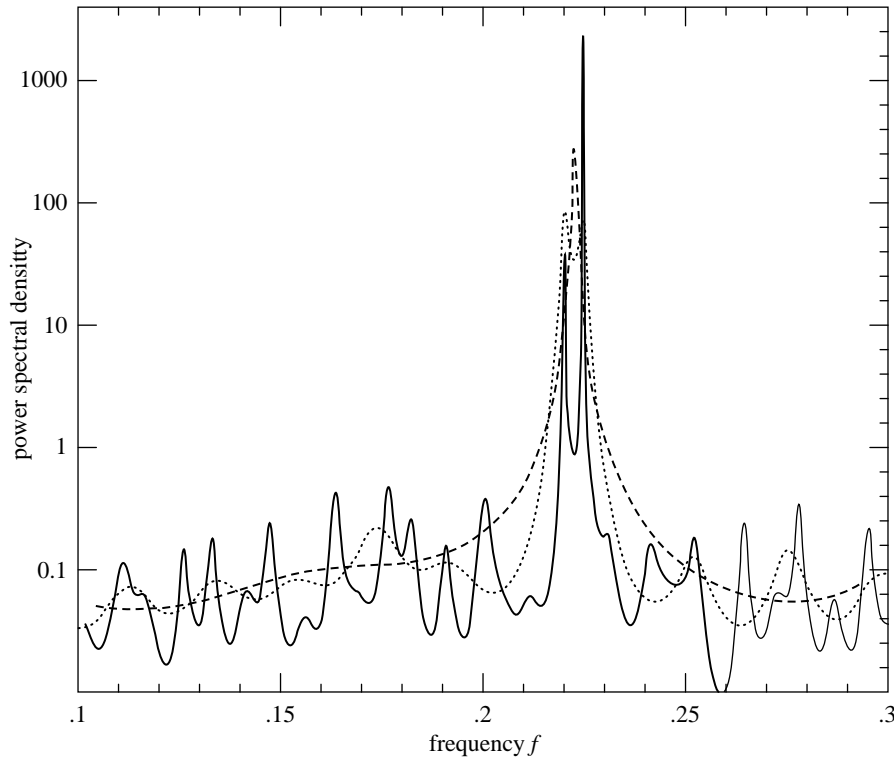


Figure 13.7.1. Sample output of maximum entropy spectral estimation. The input signal consists of 512 samples of the sum of two sinusoids of very nearly the same frequency, plus white noise with about equal power. Shown is an expanded portion of the full Nyquist frequency interval (which would extend from zero to 0.5). The dashed spectral estimate uses 20 poles; the dotted, 40; the solid, 150. With the larger number of poles, the method can resolve the distinct sinusoids; but the flat noise background is beginning to show spurious peaks. (Note logarithmic scale.)

```

sumi=0.
do ii i=1,m                Loop over the terms in the sum.
  wtemp=wr
  wr=wr*wpr-wi*wpi
  wi=wi*wpr+wtemp*wpi
  sumr=sumr-d(i)*sngl(wr)
  sumi=sumi-d(i)*sngl(wi)
enddo ii
evlmem=xms/(sumr**2+sumi**2) Equation (13.7.4).
return
END

```

Be sure to evaluate $P(f)$ on a fine enough grid to *find* any narrow features that may be there! Such narrow features, if present, can contain virtually all of the power in the data. You might also wish to know how the $P(f)$ produced by the routines `memcof` and `evlmem` is normalized with respect to the mean square value of the input data vector. The answer is

$$\int_{-1/2}^{1/2} P(f\Delta)d(f\Delta) = 2 \int_0^{1/2} P(f\Delta)d(f\Delta) = \text{mean square value of data} \quad (13.7.8)$$

Sample spectra produced by the routines `memcof` and `evlmem` are shown in Figure 13.7.1.

CITED REFERENCES AND FURTHER READING:

Childers, D.G. (ed.) 1978, *Modern Spectrum Analysis* (New York: IEEE Press), Chapter II.
 Kay, S.M., and Marple, S.L. 1981, *Proceedings of the IEEE*, vol. 69, pp. 1380–1419.

13.8 Spectral Analysis of Unevenly Sampled Data

Thus far, we have been dealing exclusively with evenly sampled data,

$$h_n = h(n\Delta) \quad n = \dots, -3, -2, -1, 0, 1, 2, 3, \dots \quad (13.8.1)$$

where Δ is the sampling interval, whose reciprocal is the sampling rate. Recall also (§12.1) the significance of the Nyquist critical frequency

$$f_c \equiv \frac{1}{2\Delta} \quad (13.8.2)$$

as codified by the sampling theorem: A sampled data set like equation (13.8.1) contains *complete* information about all spectral components in a signal $h(t)$ up to the Nyquist frequency, and scrambled or *aliased* information about any signal components at frequencies larger than the Nyquist frequency. The sampling theorem thus defines both the attractiveness, and the limitation, of any analysis of an evenly spaced data set.

There are situations, however, where evenly spaced data cannot be obtained. A common case is where instrumental drop-outs occur, so that data is obtained only on a (not consecutive integer) subset of equation (13.8.1), the so-called *missing data* problem. Another case, common in observational sciences like astronomy, is that the observer cannot completely control the time of the observations, but must simply accept a certain dictated set of t_i 's.

There are some obvious ways to get from unevenly spaced t_i 's to evenly spaced ones, as in equation (13.8.1). Interpolation is one way: lay down a grid of evenly spaced times on your data and interpolate values onto that grid; then use FFT methods. In the missing data problem, you only have to interpolate on missing data points. If a lot of consecutive points are missing, you might as well just set them to zero, or perhaps “clamp” the value at the last measured point. However, the experience of practitioners of such interpolation techniques *is not reassuring*. Generally speaking, such techniques perform poorly. Long gaps in the data, for example, often produce a spurious bulge of power at low frequencies (wavelengths comparable to gaps).

A completely different method of spectral analysis for unevenly sampled data, one that mitigates these difficulties and has some other very desirable properties, was developed by Lomb [1], based in part on earlier work by Barning [2] and Vaníček [3], and additionally elaborated by Scargle [4]. The Lomb method (as we will call it) evaluates data, and sines and cosines, only at times t_i that are actually measured. Suppose that there are N data points $h_i \equiv h(t_i)$, $i = 1, \dots, N$. Then first find the mean and variance of the data by the usual formulas,

$$\bar{h} \equiv \frac{1}{N} \sum_1^N h_i \quad \sigma^2 \equiv \frac{1}{N-1} \sum_1^N (h_i - \bar{h})^2 \quad (13.8.3)$$

Now, the Lomb *normalized periodogram* (spectral power as a function of angular frequency $\omega \equiv 2\pi f > 0$) is defined by

$$P_N(\omega) \equiv \frac{1}{2\sigma^2} \left\{ \frac{\left[\sum_j (h_j - \bar{h}) \cos \omega(t_j - \tau) \right]^2}{\sum_j \cos^2 \omega(t_j - \tau)} + \frac{\left[\sum_j (h_j - \bar{h}) \sin \omega(t_j - \tau) \right]^2}{\sum_j \sin^2 \omega(t_j - \tau)} \right\} \quad (13.8.4)$$

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